



Child maltreatment increases sensitivity to adverse social contexts: Neighborhood physical disorder and incident binge drinking in Detroit

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ABSTRACT

Introduction: Exposure to child maltreatment is associated with elevated risk for behavioral disorders in adulthood. One explanation for this life-course association is that child maltreatment increases vulnerability to the effects of subsequent stressors; however, the extent to which maltreatment increases sensitivity to social context has never been examined. We evaluated whether the association between neighborhood physical disorder and binge drinking was modified by child maltreatment exposure.

Methods: Data were drawn from the Detroit Neighborhood Health Study, a prospective representative sample of predominately African Americans in the Detroit population. Neighborhood physical disorder was measured via systematic neighborhood assessment. Child maltreatment indicators included self-reported physical, sexual, and emotional abuse. Incident binge drinking was defined as at least one episode of ≥ 5 drinks (men) or ≥ 4 drinks (women) in the past 30-day period among those with no binge drinking at baseline ($N = 1013$).

Results: Child maltreatment and neighborhood physical disorder interacted to predict incident binge drinking ($B = 0.16$, $p = 0.02$) and maximum number of past 30-day drinks ($B = 0.15$, $p = 0.04$), such that neighborhood physical disorder predicted problematic alcohol use only among individuals with high exposure to child maltreatment.

Conclusion: The results add to the growing literature that African Americans in the US are exposed to an array of stressors that have pernicious consequences for problematic alcohol use. Our results document the need for increased attention to the potential for at-risk alcohol use among populations with a high degree of stress exposure.

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1. Introduction

The World Health Organization has called for the development of global strategies to reduce alcohol consumption (Casswell and Thamarangsi, 2009), noting the high social, medical, and economic cost of excessive alcohol consumption worldwide (Jarl et al., 2008; Rehm et al., 2009). Excessive alcohol use is the third leading cause of mortality in the US (Mokdad et al., 2004), and binge drinking, defined as five or more drinks for men and four or more drinks for women per drinking occasion (National Institute on Alcohol Abuse and Alcoholism, 2005) is associated with numerous adverse health outcomes. Acute effects of binge drinking include intentional as well as unintentional injury (Hingson et al., 2002), intimate partner violence (Thompson and Kingree, 2006), unintended pregnancy

(Naimi et al., 2003a), and fetal alcohol syndrome (May and Gossage, 2001), and chronic binge drinking is associated with several forms of cancer (Longnecker et al., 1994), disruption in liver function, and premature mortality (Holman et al., 1996). Binge drinking is implicated in over half of the alcohol-attributable deaths in the US (Naimi et al., 2003b).

At the individual level, risk factors for binge drinking include young adulthood, White race/ethnicity, male gender, and psychiatric disorders (Hasin et al., 2007). Substantial evidence indicates that individuals with a history of maltreatment in childhood are at higher risk for at-risk drinking later in life (Hussey et al., 2006; Kessler et al., 1997; MacMillan et al., 2001; Widom et al., 1995). Population-based prospective data have documented high rates of binge drinking and alcohol abuse/dependence among adults reporting experiences of physical and emotional abuse (MacMillan et al., 2001), neglect (Hussey et al., 2006; Widom et al., 1995, 2007), and sexual abuse (Kessler et al., 1997; MacMillan et al., 2001; Molnar et al., 2001) in childhood.

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In addition to these individual-level stressors, a growing literature suggests that residence in disorganized and chaotic neighborhoods can be a stressful experience with deleterious implications for mental health, including substance use (Galea et al., 2004; Gmel et al., 2004). Elevated rates of alcohol use (Cerdeira et al., 2010; Galea et al., 2007; Rice et al., 1998) and binge drinking (Bernstein et al., 2007; Cerdeira et al., 2010; Hill and Angel, 2005) have been documented consistently among individuals living in neighborhoods characterized by poverty, inequality and physical disorder (with a few notable exceptions (Ennett et al., 1997; Pollack et al., 2005)). Neighborhoods with lower social cohesion have higher rates of alcohol use disorders (Winstanley et al., 2008) and alcohol-related arrests (Duncan et al., 2002) than high-cohesion neighborhoods. Further, disorganized and chaotic neighborhoods have higher alcohol outlet density (Gruenewald et al., 2002; Scribner et al., 1994), more permissive social norms with respect to drinking and drunkenness (Ahern et al., 2008), making the use of alcohol a more accessible and socially sanctioned coping strategy to manage the chronic stress of living in an unsafe and disordered area (Boardman et al., 2001). In summary, exposure to neighborhood physical disorder is a chronic stressful experience associated with elevated use of alcohol, potentially as a coping strategy to manage stress. However, not all individuals who confront stressful living conditions demonstrate elevated levels of alcohol consumption. Yet we are unaware of previous research examining factors that promote alcohol consumption in the presence of neighborhood physical disorder.

The experience of child maltreatment may be one such effect-modifying factor. Accumulating evidence indicates that child maltreatment increases vulnerability to the deleterious mental health effects of stressors that occur later in development, which has been referred to as “stress sensitization” (Hammen et al., 2000; McLaughlin et al., 2010a). The association between adult stressful life events and risk for mood and anxiety disorders is stronger among individuals with high exposure to child maltreatment (Espejo et al., 2007; Hammen et al., 2000; Kendler et al., 2004; McLaughlin et al., 2010a). Dysregulation in physiological stress response systems is thought to underlie this increased stress sensitivity (Essex et al., 2002; Heim and Nemeroff, 2001; Loman and Gunnar, 2010; Oosterman et al., 2010). Child maltreatment is also associated with increased emotional reactivity (Glaser et al., 2006; McLaughlin et al., 2010b; Wichers et al., 2009) as well as disruptions in the ability to adaptively modulate negative emotions (McLaughlin and Hatzenbuehler, 2009; McLaughlin et al., 2009). These patterns of emotional responses among maltreated individuals may increase liability to alcohol misuse in the context of adult stress. Specifically, individuals who are emotionally reactive and who have limited skills for effectively modulating their emotions may turn to alcohol or other substances to manage negative affect and arousal following stress (Carpenter and Hasin, 1998; Cooper et al., 1995; Ham and Hope, 2003). To date, however, stress sensitization effects have been examined almost exclusively in relation to mood and anxiety disorders (Espejo et al., 2007; Hammen et al., 2000; Kendler et al., 2004; McLaughlin et al., 2010a) with the exception of one study showing that adult stressors were more likely to precipitate acts of interpersonal violence among individuals with high levels of childhood maltreatment (Roberts et al., 2011).

The current study addresses two notable gaps in the literature. First, it is unclear how individual-level characteristics increase vulnerability to binge drinking among those exposed to the chronic stress of living in a disordered neighborhood. Second, the stress sensitization hypothesis has never been examined in predicting at-risk alcohol use, a prevalent and often debilitating health behavior. We address these gaps by examining whether exposure to childhood maltreatment increases susceptibility to at-risk drinking in the context of disordered neighborhood environments in

adulthood in a prospective community-based sample of predominantly African American individuals living in Detroit, Michigan. Specifically, we assess whether the relation between neighborhood physical disorder and binge drinking differs according to prior exposure to child maltreatment. Our specific aims are as follows: first, we examine whether neighborhood physical disorder, assessed using a standardized objective measure completed by independent raters, predicts incident binge drinking over the course of one year. Second, we examine whether reports of childhood maltreatment moderate the effect of neighborhood physical disorder on binge drinking. We include rigorous controls not only for established individual-level risk factors such as age, sex, and indicators of socio-economic position, but also for contextual-level confounders. Further, we examine evidence for a similar effect modifying effect for an alternative measure of at-risk drinking: largest number of drinks in the past 12 months. Given documented associations between neighborhood socio-economic position and binge drinking (Jones-Webb et al., 1997; Pollack et al., 2005), we control for average neighborhood income, house price, and percent living below the poverty line. Based on previous research documenting the sensitizing effects of childhood maltreatment on risk for psychopathology following stressful life events in adulthood, we predict that the association of neighborhood physical disorder with incident binge drinking is stronger among individuals with high exposure to child maltreatment than among individuals with low exposure to maltreatment.

2. Methods

2.1. Study sample

Data were drawn from the Detroit Neighborhood Health Study (DNHS), a longitudinal cohort of predominately African American adults (18+) living in Detroit, Michigan. Wave 1 was conducted between 2008 and 2009. Participants were selected using a dual-frame probability design, using telephone numbers obtained from the US Postal Service Delivery Sequence Files as well as a list-assisted random-digit-dial frame. Individuals without listed landlines or telephones and individuals with only a cell phone listed were invited to participate through a postal mail effort. A total of 1547 individuals participated in Wave 1, with an overall participation rate among eligible persons of 53%. Each survey participant was assigned weights that accounted for the method of obtaining contact information and the selection probability of households of different sizes and with a different number of telephone lines, and adjusted the sample to be representative of the Detroit population on a wide range of socio-demographic characteristics. Further information regarding baseline sampling can be found elsewhere (Goldmann et al., in press; Koenen et al., 2011; Uddin et al., 2010). Wave 2 was conducted one year following Wave 1; a total of 1054 individuals were re-interviewed (68% of the baseline sample). Those who did not respond were younger ($\chi^2 = 8.8, p < 0.01$), less educated ($\chi^2 = 19.0, p < 0.01$), more likely to be unemployed ($\chi^2 = 9.3, p = 0.01$) and single ($\chi^2 = 31.6, p < 0.01$); sample weights were incorporated to account for informative attrition between Waves 1 and 2. The present study focuses on the subset of individuals who reported no history of binge drinking at baseline in order to capture incident cases ($N = 1013$); sensitivity analyses were performed on the whole sample.

Participants at both waves completed a 40-min telephone survey. Informed consent was obtained at the beginning of each interview, and respondents were offered \$25 for their participation in each interview. The Institutional Review Board of the University of Michigan reviewed and approved the study protocol.

2.2. Measures

2.2.1. Alcohol consumption. Two alcohol outcomes were assessed. *Maximum drinks.* Participants reported the largest number of drinks consumed on days they drank in the 30-days prior to the interview (range 0–20). *Binge drinking.* Using the maximum drinks variable, binge drinking was defined as consuming five or more drinks (men) or four or more drinks (women) at least once in the 30-days prior to the interview.

2.2.2. Neighborhood physical disorder. Systematic assessments of neighborhood physical environment in a random sample of 138 block groups in all 54 Detroit neighborhoods were conducted by trained independent observers at Wave 1. Observers rated the neighborhood physical environment on 19 items adapted from the New York City IMPACT neighborhood evaluation scale (Ompad et al., 2008). Factor analyses suggested that three items formed a cohesive measure of the external physical environment: (1) presence of buildings with broken windows, boarded up windows, or boarded up doors; (2) presence of buildings with outside damage that

Table 1
Socio-demographic characteristics comparing those with incident binge drinking to those without in a prospective community sample of individuals in Detroit, Michigan (N = 1013).

	N	Incident binge drinking, % (SE) (N = 26)	No incident binge drinking, % (SE) (N = 987)	Chi-square, df, p-value
Sex				
Male	401	72.1 (10.6)	45.6 (2.7)	3.7, 1, 0.04
Female	612	27.9 (10.6)	54.4 (2.7)	
Age				
18–34	125	24.6 (9.1)	32 (2.9)	2.6, 5, 0.02
35–44	154	4.8 (4.7)	14.9 (1.7)	
45–54	227	35 (13.8)	23.1 (2.1)	
55–64	258	32.9 (11.9)	15.6 (1.5)	
65–74	139	2.7 (2.0)	9 (1.2)	
75+	104	0 (0.0)	5.4 (0.8)	
Race/ethnicity				
Black	863	89.9 (7.1)	88.4 (1.7)	1.4, 2, 0.24
White	111	9.1 (7.0)	7.6 (1.3)	
Other	39	1.0 (0.9)	4.0 (1.2)	
Marital status				
Married	272	17.8 (9.3)	29.1 (2.2)	1.1, 2, 0.34
Divorced/separated/widowed	403	46.8 (13.1)	26.1 (2.0)	
Never married	338	35.4 (12.1)	44.8 (2.7)	
Education				
Less than high school	128	6.4 (5.2)	15.5 (2.0)	1.4, 2, 0.25
High school or GED	288	59.3 (12.9)	42.3 (2.7)	
More than high school	597	34.3 (12.7)	42.2 (2.5)	
Personal income				
Less than 25,000	276	29.1 (11.5)	33.6 (2.7)	0.2, 2, 0.85
25,000–50,000	241	33.7 (13.1)	26.4 (2.4)	
50,000+	354	37.2 (14.4)	40.1 (2.7)	
Employment				
Looking for work or unemployed	156	19.8 (9.0)	25.9 (2.7)	0.4, 1, 0.53
Other	857	80.2 (9.0)	74.1 (2.7)	

can only be corrected by major repairs such as siding, shingles, boards, brick, concrete, and stucco; and (3) presence of entirely vacant buildings. These measures also demonstrated excellent internal consistency ($\alpha = 0.83$). The factor analyses yielded principal component values for each block group, which were then tested for and showed spatial autocorrelation in ArcMap Version 9.2 (ESRI, Redlands, CA) (Moran's $I = 0.14$, z -score = 4.08, $p < 0.01$). We then used ordinary kriging methods (Koizumi et al., 2009; Kumar et al., 2007) within ArcMap Version 9.2 to predict frequency values for those areas of Detroit that were not evaluated and created maps of these prediction estimates. Predicted values were then averaged by neighborhood. We used these values to create a continuous score representing the physical organization of each respondent neighborhood (mean = -0.12 , range -1.52 to 1.23), and then categorized the variable into quartiles to detect potential non-linear relationships with alcohol outcomes. Because our main effect analyses indicated that neighborhood disorganization was monotonically associated with alcohol outcomes, we used a continuous variable for analyses of interaction.

2.2.3. Childhood maltreatment. Childhood maltreatment questions were included at Wave 2 and drawn from the conflict tactics scale (CTS) (Straus, 1979) and the Childhood Trauma Questionnaire (CTQ) (Bernstein et al., 1997). CTS items assessed physical abuse (e.g., "People in my family hit me so hard that it left me with bruises and marks"), and emotional abuse (e.g., "People in my family said hurtful or insulting things to me") assessed before age 11. Response options were rated on a five point scale and ranged from "never true" (1) to "very often true" (5). CTQ items assessed physical abuse before age 18 (e.g., "Did your parent, step-parent, or adult guardian ever push, grab, or shove you") with response options ranging from "never" to "more than a few times" and sexual abuse before age 18 (e.g., "Were you ever touched in a sexual way by an adult or older child...") which was coded as occurring once, multiple times, or never.

CTS physical abuse and emotional abuse questions as well as CTQ physical and sexual abuse questions were recoded into three-level variables indicating whether each abuse type occurred: (1) never, (2) rarely or sometimes, or (3) often. Scores were summed to create a continuous variable ranging from 0 to 22 which was normally distributed in the sample. We also dichotomized the child maltreatment variable at the 75th percentile.

2.2.4. Control variables. Individual-level control variables included age, sex, race/ethnicity, income, education, employment status, and marital status. Variable coding is shown in Table 1. As described above, we controlled for three indicators of neighborhood socio-economic position to document the specific effect of neighborhood physical disorder. Neighborhood-level covariates included three variables assessed at the level of Census tracts (US Bureau of the Census, 2000): percent living below the poverty line, median house value, and median income. These variables

were categorized into quartiles. Preliminary analyses indicated that neighborhood disorder measured at the neighborhood level was significantly associated with each Census tract socio-economic indicator at $p < 0.0001$.

2.3. Statistical analysis

All analyses were conducted using SAS-callable SUDAAN software (Research Triangle Institute, 2009), and adjusted for the complex survey design as well as attrition weights. Bivariate associations between exposure variables and binge drinking were examined using cross-tabulations, chi-square tests, and multivariate associations were examined using logistic regression. Bivariate and multivariate associations between exposure variables and maximum number of drinks were assessed with Poisson regression due to non-normality of the distribution. Interactions between child maltreatment and neighborhood physical disorder were assessed on the multiplicative scale. Preliminary analyses indicated that continuous measures of childhood maltreatment and neighborhood physical disorder were uncorrelated (Pearson's correlation coefficient = 0.04, $p = 0.16$). Statistical significance was evaluated using 2-sided 0.05 level tests.

3. Results

The incidence of binge drinking was 3.48% (SE = 0.9). Binge drinking was more common among males ($\chi^2 = 3.7$, $df = 1$, $p = 0.04$) and those in middle age (45–64; $\chi^2 = 2.6$, $df = 5$, $p = 0.02$) (Table 1). Examining socio-demographic associations with neighborhood physical disorder, those living in disordered neighborhoods were more likely to be never married ($\chi^2 = 5.4$, $df = 2$, $p = 0.005$) and make less than \$25,000 in the last year ($\chi^2 = 4.3$, $df = 2$, $p = 0.01$).

Table 2 shows the main effects of childhood maltreatment and neighborhood physical disorder on binge drinking and maximum drinks. First we examined a continuous variable summing all maltreatment experiences; childhood maltreatment was significantly associated with binge drinking. Each one-unit increase in the childhood maltreatment score was associated with 1.1 times the odds of binge drinking (95% C.I. 1.01–1.23). When maltreatment types were examined individually, sexual abuse was associated with 3.1 times increased odds of binge drinking (95% C.I. 1.04–9.35). The

Table 2
Main effects of childhood maltreatment and neighborhood physical disorder on incident binge drinking and maximum drinks in the past 30-days among respondents with no binge drinking at the baseline interview in a prospective community sample of individuals in Detroit, Michigan ($N = 1013$).

	<i>N</i>	% Incident binge drinking	OR (95% C.I.)	OR ^a (95% C.I.)	Mean maximum drinks in the past 30 days	Unadjusted incidence density ratio	Adjusted incidence density ratio ^a
Childhood maltreatment^a							
Cumulative childhood maltreatment		–	1.10 (1.01–1.20)	1.11 (1.01–1.23)	–	1.02 (0.99–1.05)	1.03 (0.99–1.06)
Physical abuse	277	5.42	1.78 (0.58–5.40)	2.59 (0.83–8.11)	2.58	1.18 (0.95–1.48)	1.2 (0.95–1.51)
No physical abuse	736	3.12	1	1	2.18		1
Emotional abuse	171	5.69	1.99 (0.65–6.09)	1.00 (0.34–2.97)	2.75	1.24 (0.96–1.60)	1.18 (0.97–1.42)
No emotional abuse	842	2.94	1	1	2.22		1
Sexual abuse	188	6.56	3.26 (1.18–9.01)	3.12 (1.04–9.35)	2.51	1.09 (0.81–1.46)	1.22 (0.92–1.62)
No sexual abuse	825	2.11	1	1	2.31		1
Neighborhood physical disorder^b							
Lowest quartile (most disordered)	257	4.52	1.14 (0.31–4.27)	1.72 (0.40–7.44)	2.76	1.28 (0.90–1.81)	1.28 (0.95–1.72)
2nd quartile	284	2.94	0.68 (0.17–2.70)	1.22 (0.27–5.48)	2.34	1.08 (0.81–1.44)	1.10 (0.84–1.44)
3rd quartile	219	2.74	0.73 (0.16–3.35)	1.15 (0.28–4.65)	2.28	1.05 (0.79–1.40)	1.10 (0.86–1.41)
Highest quartile (least disordered)	253	3.97	1	1	2.16	1	1

^a Controlled for age, sex, race, unemployment, education, and income.

^b Higher score is less disordered. Continuous factor score representing three items: presence of buildings with broken windows, boarded up windows, or boarded up doors; presence of buildings with outside damage that can only be corrected by major repairs such as siding, shingles, boards, brick, concrete, and stucco; and presence of entirely vacant buildings.

^c Physical abuse coded positive if greater than 7; emotional neglect coded positive if >2; sexual abuse coded positive if >0; cumulative childhood maltreatment entered as continuous variable.

associations between neighborhood physical disorder and alcohol outcomes were not statistically significant.

Table 3 shows the interaction of neighborhood physical disorder and childhood maltreatment predicting binge drinking and maximum number of drinks. In regression models adjusted for demographics and indicators of neighborhood socio-economic position, a significant interaction between neighborhood physical disorder and childhood maltreatment was observed in predicting both binge drinking ($\beta = 0.15$, $SE = 0.07$, $p = 0.04$) and maximum number of drinks ($\beta = 0.05$, $SE = 0.02$, $p = 0.01$). As a check on the robustness of the results, we reran these models with prevalent binge drinking (any binge drinking reported at Wave 1 or 2) as the outcome. We continued to observe a significant interaction

($\beta = 0.26$, $SE = 0.04$, $p < 0.01$), indicating that our results are not limited only to the incident binge drinkers in this sample.

To better understand the direction and magnitude of the interaction, in Fig. 1 we stratified by the dichotomized childhood maltreatment variable (dichotomized at the 75th percentile) and examined the associations of the dichotomous neighborhood physical disorder variable with incident binge drinking. The incidence of binge drinking was 12.2% among those with high exposure to childhood maltreatment in the highest quartile of neighborhood physical disorder, compared to 1.74% among those with high childhood maltreatment exposure but in the lowest quartile of neighborhood physical disorder and 1.71% among those in the highest quartile of neighborhood physical disorder but below the

Table 3
Effect measure modification between neighborhood physical disorder and childhood maltreatment in predicting Wave 2 binge drinking and maximum number of drinks in the past 30 days among those with no binge drinking at baseline in a prospective community sample of individuals in Detroit, Michigan ($N = 1013$).

	Beta, SE, <i>p</i> -value		
	Model 1: unadjusted	Model 2: adjusted for demographics ^c	Model 3: adjusted for demographics and other neighborhood characteristics ^c
Binge drinking at Wave 2			
Neighborhood physical disorder ^a	–1.33 (SE = 0.60), $p = 0.03$	–1.33 (SE = 0.60), $p = 0.03$	–1.21 (SE = 0.55), $p = 0.03$
Childhood maltreatment ^b	0.06 (SE = 0.06), $p = 0.29$	0.06 (SE = 0.06), $p = 0.29$	0.05 (SE = 0.06), $p = 0.43$
Interaction between neighborhood disorder and childhood maltreatment	0.16 (SE = 0.07), $p = 0.02$	0.16 (SE = 0.07), $p = 0.02$	0.15 (SE = 0.07), $p = 0.04$
Maximum number of drinks at Wave 2			
Neighborhood physical disorder ^a	–0.29 (SE = 0.12), $p = 0.01$	–0.71 (SE = 0.19), $p < 0.001$	–0.90 (SE = 0.34), $p < 0.01$
Childhood maltreatment	0.01 (SE = 0.01), $p = 0.25$	–0.01 (SE = 0.01), $p = 0.21$	–0.02 (SE = 0.01), $p = 0.10$
Interaction between neighborhood disorder and childhood maltreatment	0.03 (SE = 0.02), $p = 0.11$	0.05 (SE = 0.02), $p = 0.004$	0.05 (SE = 0.02), $p = 0.01$

^a Higher score is less disordered. Continuous factor score representing three items: presence of buildings with broken windows, boarded up windows, or boarded up doors; presence of buildings with outside damage that can only be corrected by major repairs such as siding, shingles, boards, brick, concrete, and stucco; and presence of entirely vacant buildings.

^b Continuous variable with range 0–22.

^c Model 2 was adjusted for age, sex, race/ethnicity, income, education, employment status, and marital status. Model three was adjusted for aforementioned variables and additionally controlled for the percent in the census tract that were living below the poverty line, the median house value of the census tract, and the median income of the census tract.

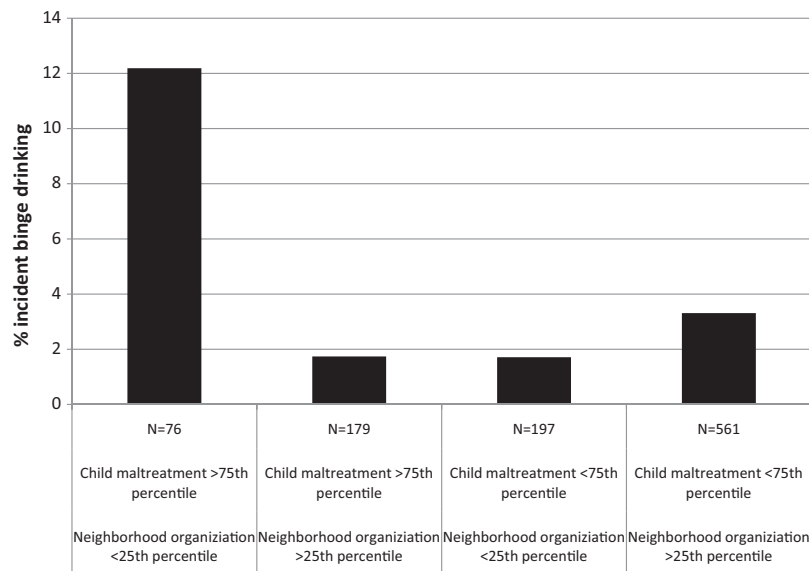


Fig. 1. Incidence of binge drinking at Wave 2 among those low versus high on neighborhood physical disorder and childhood maltreatment exposure among those with no binge drinking at baseline in a prospective community sample of individuals in Detroit, Michigan ($N = 1013$).

75th percentile on childhood maltreatment. Results were similar for maximum number of drinks (not shown): the mean maximum drinks was 2.74 among those with high exposure to childhood maltreatment in the highest quartile of neighborhood physical disorder, compared to 2.21 among those with high childhood maltreatment exposure but in the lowest quartile of neighborhood physical disorder and 1.85 among those in the highest quartile of neighborhood physical disorder but below the 75th percentile childhood maltreatment.

4. Discussion

Exposure to childhood maltreatment modifies the effect of neighborhood physical disorder on risk of binge drinking among this predominately African American community sample in Detroit, MI. Specifically, high levels of neighborhood physical disorder were associated with elevated risk of incident binge drinking, but only among individuals with high exposure to maltreatment in childhood. The interaction of child maltreatment and neighborhood physical disorder predicted alcohol outcomes even after controlling for individual-level risk factors and neighborhood socio-economic indicators, indicating that the physical characteristics of a neighborhood increases risk of binge drinking, but only among those with a pre-existing vulnerability conferred by early-life exposure to maltreatment. These findings are consistent with the theory that stressors experienced in childhood increase vulnerability to the deleterious consequences of stressful life circumstances in adulthood, thereby heightening risk for adverse health consequences in the context of adult stress (Hammen et al., 2000; Kendler et al., 2004; Espejo et al., 2007; McLaughlin et al., 2010a,b).

We document that the one-year incidence of binge drinking in this sample was 2.56%, which is similar to an estimated national one-year incidence of 3.18%, calculated using available data from the follow-up of the National Epidemiologic Survey on Alcohol and Related Conditions (NESARC). However, we note that previous studies have documented that the effect of neighborhood poverty on alcohol use is stronger for African Americans compared with Whites (Jones-Webb et al., 1997). Because our sample is predominately African American and representative of the Detroit community, these results may not generalize to other populations with a different distribution of neighborhood factors and other

stressors. However, the results add to the growing literature that African Americans in the US are exposed to an array of stressors that have pernicious consequences for problematic alcohol use (Keyes et al., 2011a), despite exhibiting lower rates of alcohol consumption and alcohol use disorders compared with Whites (Hasin et al., 2007). Our results document the need for increased attention to the potential for problematic alcohol use among populations with a high degree of stress exposure.

Our findings are innovative in several respects. To our knowledge, this is the first study to explore the stress sensitization hypothesis using a neighborhood-level exposure. Our results suggest that the stressful experience of living in a neighborhood characterized by physical disorder increases risk for adverse health outcomes among those with high exposure to childhood maltreatment. A similar sensitizing effect of exposure to child maltreatment has been found in predicting mental health outcomes such as mood and anxiety disorders (Espejo et al., 2007; Hammen et al., 2000; Kendler et al., 2004; McLaughlin et al., 2010a), and preliminary evidence suggests a similar sensitizing role of childhood maltreatment in predicting perpetration of interpersonal violence following adult stress exposure (Roberts et al., 2011). This prior research, however, has evaluated adult stressful life events including interpersonal, occupational, financial, and legal events, and no previous literature has documented a stress sensitization effect for alcohol outcomes. Our findings add to a growing body of research supporting the hypothesis that the negative health consequences of stressful experiences in adulthood, including those conferred by social and contextual factors, are heightened by exposure to childhood maltreatment. These findings suggest that exposure to early-life stress creates a diathesis for psychopathology and substance misuse that can be elicited by stress exposure later in life. Finally, our findings support the hypothesis that individual behavior can be influenced by deleterious features of the social environment; when potentiated by exposures endured throughout the life course, this cascade of stress exposure results in adverse mental health outcomes. We also document main effects of childhood maltreatment on at-risk drinking. Specifically, we found that sexual abuse significantly predicted incident binge drinking, consistent with substantial epidemiologic evidence documenting the pernicious consequences of sexual abuse for substance use outcomes (see review in Keyes et al., 2011b). Further, these findings

provide additional support for neighborhood physical disorder as a risk factor for adverse health outcomes in the presence of childhood maltreatment, independent of the neighborhood socio-economic position. Although studies have been mixed on the association between neighborhood physical disorder and crime rates (a hypothesis known as ‘broken windows’ theory) (Cerdeira et al., 2009), physical disorder is consistently associated with health outcomes such as sexually transmitted infections (Cohen et al., 2003), mental health (Aneshensel and Sucoff, 1996; Kruger et al., 2007), and obesity (Franzini et al., 2009). Our analyses extend previous studies of the neighborhood context by using independent raters rather than inhabitant self-report to characterize the neighborhood environment. Further, we used rigorous controls for neighborhood socio-economic factors, indicating that physical disorder of the neighborhood is associated with elevated risk for the onset of binge drinking patterns over and above the effects of neighborhood poverty and disadvantage. An important next step in this research program is to determine whether these results generalize to other alcohol-related health outcomes, including alcohol use disorders, liver function, and chronicity of heavy drinking. Mechanisms for this effect may be the increase in alcohol outlets among disorganized neighborhoods (Gruenewald et al., 2002; Scribner et al., 1994), fewer social restrictions on drinking and drunkenness (Ahern et al., 2008), or psychological stress generated by efforts to cope with a negative physical environment (Bernstein et al., 2007). However, we also note that individuals who consume alcohol or possess risk factors associated with heavy alcohol consumption may be more likely to live in or move to neighborhoods that are more disorganized; thus, the causal effect of these associations may be reversed. Although these data are consistent with the hypothesis that neighborhood disorganization is a risk factor for binge drinking in the presence of childhood maltreatment, a social selection hypothesis cannot be ruled out.

These results also add to a growing body of literature marking the importance of studying cumulative adversities experienced across the life course to understand adult health outcomes such as binge drinking. Previous studies have shown that the trajectory of socio-economic status from childhood to adulthood influences alcohol-related and other health risks (Fox et al., 2010; Lloyd and Turner, 2008; Pensola and Martikainen, 2003; Pensola and Valkonen, 2000), especially downward mobility (Pensola and Valkonen, 2000), further indicating that adversity is a life course process and that risks interact to increase the risk for adverse health outcomes. Research is also accumulating to document genetic factors that interact with childhood maltreatment experiences in impacting alcohol use and other externalizing behaviors from adolescence to adulthood (Caspi et al., 2002; Enoch et al., 2010; Kaufman et al., 2007; Kim-Cohen et al., 2006; Widom and Brzustowicz, 2006). These same genetic factors also interact with the social environment to raise risk for psychopathology (Koenen et al., 2009), suggesting that the interplay of stressors across the life course and genetic vulnerability provide a compelling pathway for the incidence and persistence of adverse mental health outcomes.

Because both child maltreatment and neighborhood physical disorder are associated with other adverse exposures that increase the risk for problematic alcohol use, our findings highlight just one piece of a complex set of pathways that link the social environment to patterns of alcohol use in adulthood. We note that childhood maltreatment is correlated with a wide range of adversities in childhood including instability in the home (Dube et al., 2002; Dong et al., 2004), poverty (Turner et al., 2006), and parental dysfunction across multiple domains including substance abuse, criminality, and psychopathology (Dinwiddie and Bucholz, 1993; Kelleher et al., 1994; Chaffin et al., 1996; Conron et al., 2009). Further, maltreatment in childhood raises the risk for other factors throughout the life course that may mediate the associations observed here, including

both perpetrating and being the victim of interpersonal violence in adulthood (Desai et al., 2002; Ehrensaft et al., 2003; McKinney et al., 2009) and conduct disorder in adolescence and antisocial personality disorder in adulthood (Iacono et al., 1999; Jessor and Jessor, 1977). These factors were unmeasured in the present data, thus future research should probe these associations further for specificity across multiple domains of childhood maltreatment and mediators through the life course.

Results should be interpreted with study limitations in mind. Childhood maltreatment is retrospectively self-reported by respondents, which potentially introduces bias into the results (Green et al., 2010; Hardt and Rutter, 2004). Longitudinal studies have documented that child abuse reports, including childhood sexual abuse, are unstable over time (Widom, 1996, 1997). Studies testing the validity of retrospective self-report generally find few false positives, indicating that the consequence of unstable reporting is that the prevalence of childhood maltreatment assessed at only one time point is likely underestimated (Fergusson et al., 2000; Widom, 1996, 1997). Further, several independent cohort studies with both prospective and retrospective assessment of childhood maltreatments have reported that risk association estimates are stable regardless of the timing of assessment (Fergusson et al., 2000; Hardt et al., 2010). Thus, although the estimates provided in these data should be interpreted with caution, the associations reported here are unlikely to be artifacts of assessment timing. Additionally, neighborhood physical disorder is based on values for some Detroit block groups predicted using the values of the 138 block groups that we assessed. It is possible, though unlikely, that these values do not accurately describe the level of disorder in these block groups. However, there was strong evidence of spatial clustering of neighborhood physical disorder, which gives us confidence in our ability to using kriging techniques to predict block group values. The block groups randomly selected for evaluation were also chosen in such a way as to represent various different types of areas (based on population characteristics and density). Additionally, sensitivity analyses using only the observed neighborhood physical disorder values (rather than inclusion of predicted values) indicated that the magnitude and direction of interaction effects were similar. Further, neighborhood-level control variables were measured at the census tract level, which potentially comprises several block groups. Thus, we may have residual confounding if there is block-group heterogeneity of socio-economic indicators within the census tract group. However, all census tract indicators were strongly and significantly associated with block-level neighborhood disorganization, thus the impact of this potential residual confounding is likely minimal. Further, there is the possibility of neighborhood selection bias associated with childhood maltreatment, given that individuals with maltreatment histories may cluster within lower socio-economic neighborhoods. However, our analyses indicated that there was no correlation between maltreatment and neighborhood physical disorder in these data. Finally, we note that the participation rate at Wave 1 was 53%; thus, our findings are generalizable to populations who tend to respond to survey research of this kind. While we do not have information on whether participation was associated with the relevant study variables, this limitation should be kept in mind while interpreting the results.

Despite these limitations, the present study capitalized on the strengths of a longitudinally assessed community sample of individuals to document a negative synergistic effect of child maltreatment and neighborhood physical disorder in heightening risk for the onset of problematic patterns of alcohol consumption. Existing evidence increasingly suggests that neighborhood deprivation has deleterious effects on a wide range of poor health outcomes, including mental health problems such as depression (Latkin and Curry, 2003; Ross, 2000) and drug use disorders (Boardman et al.,

2001) as well as physical health outcomes including heart disease (Diez-Roux et al., 1997, 2001; Sundquist et al., 2004) and obesity (Franzini et al., 2009). Future research should examine the potential sensitizing effect of childhood maltreatment and exposure to other adversities in the association of neighborhood characteristics with these outcomes in order to better understand the public health consequences of exposure to stress over the life-course.

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Contributors

Katherine Keyes conceived of the present analysis, led the statistical analysis efforts, and wrote the initial drafts of the manuscript. Sandro Galea and Karestan Koenen designed the study, led data collection efforts, supervised the present analysis and provided critical revisions to the manuscript. Monica Uddin was involved with data collection and provided critical revisions to the manuscript. Emily Goldmann supervised field work during the data collection and was involved with data analysis and in providing critical revisions to the manuscript. Katie McLaughlin were involved with data analysis and provided critical revisions to the manuscript.

Conflict of interest

None.

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